

Math233 Note

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1 Lecture Video

1.1 Function of Random Variable

Recall:

$$X = \text{Normal}(\mu, \sigma)$$

means that X has probability density function (PDF):

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

Or:

$$P[X \leq x] = \Phi\left(\frac{x-\mu}{\sigma}\right)$$

Q: What if Y is a function of random variable?

A: The only way we can find PDF is through the cumulative distribution function(CDF):

$$\begin{aligned} F_Y[y] &= P[Y \leq y] = P[g(x) \leq y] \\ &\Rightarrow f_Y = F_Y' \end{aligned}$$

Example:

If X is a continuous random variable with probability density f_X . Let $Y = X^2$. For $Y \geq 0$,

$$\begin{aligned} F_Y(y) &= P[Y \leq y] \\ F_Y(y) &= P[X^2 \leq y] \\ F_Y(y) &= P[-\sqrt{y} \leq X \leq \sqrt{y}] \\ F_Y(y) &= F(\sqrt{y}) - F(-\sqrt{y}) \\ \frac{d}{dy}(F_Y(y) &= F(\sqrt{y}) - F(-\sqrt{y})) \\ f_Y(y) &= \frac{1}{2\sqrt{y}}[f_X(\sqrt{y}) + f_X(-\sqrt{y})] \end{aligned}$$

1.1.1 General Case

Consider a continuous random variable X with PDF f_X and g a monotonic differentiable function. (i.e, g^{-1} exists everywhere and $\frac{d}{dy}g^{-1}$ also exists)

Let $Y = g(X)$. Then:

$$\begin{aligned}P[Y \leq y] &= P[g(X) \leq y] = P[X \leq g^{-1}(y)] \\ &= \int_{-\infty}^{g^{-1}(y)} f_X(x)dx\end{aligned}$$

Let $z = g(x)$

$$\begin{aligned}x &= g^{-1}(z), dx = \frac{d}{dz}(g^{-1}(z))dz \\ P[Y \leq y] &= \int_{-\infty}^y f_X(g^{-1}(z)) \frac{d}{dz}(g^{-1}(z))dz \\ f_Y(y) &= f_X(g^{-1}(y)) \frac{d}{dy}(g^{-1}(y))\end{aligned}$$

Conclusion If X has PDF f_X and g strictly monotone and differentiable, then $Y = g(X)$ has PDF.

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \left| \frac{d}{dy}g^{-1}(y) \right| & \text{if } y = g(x) \\ 0 & \text{otherwise} \end{cases}$$

Recall $x = g^{-1}(y)$ is such that $g(x) = y$

1.2 Joint Distribution Function

The Joint CDF for two random variables X, Y is function F such that:

$$F(x, y) = P[X \leq x, Y \leq y]$$

1.2.1 Discrete Case

Joint Probability Mass Function:

$$\begin{aligned}p(x_i, y_j) &= P[X = x_i, Y = y_j] \\ &= P[\cup_j (X = x_i, Y = y_j)] \\ &= \sum_j p(x_i, y_j)\end{aligned}$$

1.3 Continuous Case

$$P[X, Y \in C] = \int \int_C f(x, y) dx dy$$

$$P[X \in (a, b), Y \in (c, d)] = \int_c^d \int_a^b f(x, y) dx dy$$

$$\begin{aligned} F(x, y) &= P[X \leq x, Y \leq y] \\ &= \int_{-\infty}^y \int_{-\infty}^x f(u, v) du dv \end{aligned}$$

We know:

$$\frac{d^2 F}{dx dy} = f$$

Lastly

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

Since we are only integrating over R^2 , this also holds for n random variables X_1, X_2, \dots, X_n

2 Class Lecture And Examples

Recall Fundamental Theorem of Calculus:

$$F(y) = \int_a^{h(y)} f(x) dx$$

$$F'(y) = f(h(y)) * h'(y)$$

Then, we can find PDF of $Y = g(X)$ with X being a continuous random variable

$$f_Y(y) = \begin{cases} f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| & \text{if } y = g(x) \\ 0 & \text{otherwise} \end{cases}$$

Example 1:

Let $X = \text{Uniform}(-1, 1)$. Then:

$$f(x) = \begin{cases} 1/2 & \text{on } (-1, 1) \\ 0 & \text{otherwise} \end{cases}$$

Let $Y = X^2$. We have:

$$P[Y \leq y] = P[X^2 \leq y] = P[-\sqrt{y} \leq X \leq \sqrt{y}]$$

$$\begin{aligned}
&= \int_{-\sqrt{y}}^{\sqrt{y}} f_x(x) dx \\
&= \int_{-\sqrt{y}}^{\sqrt{y}} \frac{1}{2} dx = \sqrt{y} \\
f_Y(y) &= \frac{d}{dy} \sqrt{y} = \frac{1}{2\sqrt{y}}
\end{aligned}$$

Example 2:

Let $X = \text{Normal}(0,1)$, $Y = \cos(X)$.
 $X \in (-\infty, \infty)$, $Y \in [-1,1]$

$$\begin{aligned}
P[Y \leq y] &= P[\cos(X) \leq y] \\
&\neq P[X \leq \arccos(y)] \\
&\neq \int_{-\infty}^{\arccos(y)} f(x) dx
\end{aligned}$$

This is the wrong domain. Since $Y = \cos(X)$ is an oscillating functions, we must identify the interval on which we can integrate over. We know X is a Normal Random Variable on $(0,1)$. Hence, we can find the interval by finding the intersection of $f(x) = \cos(y)$. We can then find the right region to integrate over. So:

$$P[Y \leq y] = P[\cos(X) \leq y] = \int_{\cos(x) \leq y} f(x) dx$$

Problem 5.22 (Self Test):

Let $U = \text{Uniform}(0,1)$, a and b are constants. $a < b$

a, Show that if $b \geq 0$ then bU is uniformly distributed on $(0,b)$ and if $b < 0$, then bU is uniformly distributed on $(b,0)$

b, Show that $a + U$ is uniformly distributed on $(a, 1 + a)$

c, What function of U is uniformly distributed on (a,b)

d, Show that $\min(U, 1 - U)$ is a uniform $(0,1/2)$ random variable

e, Show that $\max(U, 1 - U)$ is a uniform $(1/2, 1)$ random variable

Solution:

a, Consider $b > 0$:

$$\begin{aligned}
F(bU) &= P(bU \leq x) \\
&= P(U \leq \frac{x}{b}) \\
&= \int_0^{x/b} dx
\end{aligned}$$

$$\begin{aligned}
&= F\left(\frac{x}{b}\right) \\
&= \frac{x}{b} \\
\Rightarrow f_{bu}(x) &= F'\left(\frac{x}{b}\right) = \frac{1}{b}
\end{aligned}$$

The case $b < 0$ is left as exercise. b and c are also similar. d, Let $Y = \min(U, 1 - U)$

$$\begin{aligned}
F(Y) &= P[\min(U, 1 - U) \leq x] \\
&= P[U \leq x] \cup P[1 - U \leq x]
\end{aligned}$$

Since $U \cap (1-U) = \emptyset$. We have:

$$\begin{aligned}
&= \left(\int_0^x dx\right) + (1 - P[U < 1 - x]) \\
&= x + (1 - (1 - x)) \\
&= 2x \\
\Rightarrow f_u(x) &= F'(y) = 2
\end{aligned}$$

Which is the density function for $U = \text{Uniform}(0, 1/2)$. e) is similar and left as an exercise.