

Probability Theory II (Spring 2026)

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Primary reference: Varadhan, Stochastic Processes

Prerequisites: Varadhan, Probability Theory

Other references:

Norris, Advanced Probability

Stroock, Probability Theory, An Analytic View

Le Gall, Brownian Motion, Martingales, and Stochastic Calculus

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1. Continuous-time processes & martingales

From now on, (Ω, \mathcal{F}, P) denotes a probability space on which all random variables are defined.

1.1. Continuous-time processes

Defn. Given an index set Π , a stochastic process $X = \{X_t : t \in \Pi\}$ is a collection of random variables $X_t = X(t, \omega)$, $t \in \Pi$, $\omega \in \Omega$, taking values in some measurable space.

- If Π is a countable set, X is a discrete-time stochastic process.
- If $\Pi \subseteq \mathbb{R}$ is a finite or infinite interval, X is a continuous-time stochastic process.

Careful: For Π uncountable, events like

$$\{\omega : X_t(\omega) \in [a, b] \forall t \in \Pi\}$$

or $\{\omega : t \mapsto X_t(\omega) \text{ is continuous}\}$

are not measurable in general.

Soln: Build regularity into Ω .

Defn. $C[a,b]$ is the Banach space of continuous functions $f: [a,b] \rightarrow \mathbb{R}$ with norm

$$\|f\| = \sup_{t \in [a,b]} |f(t)|.$$

Defn. $D[a,b]$ is the Polish space (complete sep metric) of left-continuous functions with right limits (càdlàg functions) $f: [a,b] \rightarrow \mathbb{R}$ with metric

$$d(f,g) = \inf \left\{ r > 0 : \exists \lambda: [a,b] \rightarrow [a,b] \text{ increasing, invertible, and continuous s.t. } \sup_t |\lambda(t) - t| \leq r, \sup_t |f(t) - g(\lambda(t))| \leq r \right\}.$$

Prop. $D[a,b]$ is indeed a Polish space.

[Billingsley, Convergence of Prob. Measures, Chapter 3].

Prop. The σ -fields of Borel sets of $C[a,b]$ or $D[a,b]$ are generated by the cylinder sets, i.e., by

$$\{x \in X : \pi_t x = x(t) \in I\}, \quad I \in \mathcal{B}(\mathbb{R}), \quad t \in [a,b]$$

Sketch. Let \mathcal{F} denote the cylinder σ -field. Since $\{x \in X : x(t) \in I\}$ is a Borel set, $\mathcal{F} \subset \mathcal{B}$. Conversely,

$$\{x \in C : \|x\| \leq r\} = \bigcap_{t \in [a,b] \cap \mathbb{Q}} \pi_t^{-1}([-r, r]) \in \mathcal{F}.$$

Exercise: $X=D$

We will always use the cylinder σ -field (= Borel σ -field) on $C[a,b]$ or $D[a,b]$. The space $C[0,\infty)$ is defined by $x \in C[0,\infty)$ if $x|_{[0,T]} \in C[0,T]$ and likewise for $D[0,\infty)$ with unif. on compacts metric.

Defn. Let (Ω, \mathcal{F}) be a measurable space. Then a family $\{\mathcal{F}_t : t \geq 0\}$ of sub- σ -fields $\mathcal{F}_t \subset \mathcal{F}$ is

- a filtration if $\mathcal{F}_s \subseteq \mathcal{F}_t$ if $s \leq t$.
- a right-cont. filtration if $\mathcal{F}_t \supset \mathcal{F}_{t+} = \bigcap_{s>t} \mathcal{F}_s$

Defn. Let $(\Omega, \mathcal{F}, (\mathcal{F}_t))$ be a filtered measurable space. Then a stochastic process (X_t) is

- adapted if $\forall t: X_t \in \mathcal{F}_t$, i.e., X_t is \mathcal{F}_t -meas.
- integrable if $\forall t: X_t \in L^1$, i.e., $E|X_t| < \infty$.

Example. If Ω is a space of functions on $\mathbb{T} \subset [0,\infty)$,

$$\mathcal{F}_t = \sigma(\omega(s) : s \leq t)$$

is the canonical filtration. The canonical process $X_t(\omega) = \omega(t)$, $\omega \in \Omega$ is (\mathcal{F}_t) -adapted.

If Ω is a space of right-cont. functions, it is also adapted to (\mathcal{F}_{t+}) .

Prop. Let (X_t) be a $D[0, \infty)$ -valued adapted process. Then (X_t) is progressively measurable: $\forall T > 0$

$$(\omega, t) \mapsto X_t(\omega)$$

on $\Omega \times [0, T]$ is measurable w.r.t. $\mathcal{F}_T \otimes \mathcal{B}([0, T])$.

Sketch. By right-cont. can approximate by

$$X_t^n(\omega) = X_{2^{-n} \lceil 2^n t \rceil}(\omega)$$

For every $A \in \mathcal{B}(\mathbb{R})$,

$$\{(\omega, t) \in \Omega \times [0, T] : X_t^n(\omega) \in A\}$$

$$= \bigcup_{\substack{k \\ 2^{-n}k < T}} \{X_{2^{-n}k} \in A\} \times \underbrace{[2^{-n}(k-1), 2^{-n}k)}_{\in \mathcal{B}([0, T])}.$$

Thus X^n is $\mathcal{F}_T \otimes \mathcal{B}([0, T])$ -measurable. Take limits.

Sometimes it is also convenient to assume that the \mathcal{F}_t contain all null sets, i.e. is complete.

Defn. A filtration satisfies the usual conditions if it is right-continuous and complete (w.r.t. a given probability measure).

1.2. Functions of bounded variation

Defn. The total variation of $x: [a, b] \rightarrow \mathbb{R}$ is

$$V_x[a, b] = \sup \left\{ \sum_{i=1}^n |x(t_i) - x(t_{i-1})| : \underbrace{a \leq t_0 < \dots < t_n \leq b}_{\text{partition of } [a, b]} \right\}.$$

The space $BV[a, b]$ of functions of bounded variation consists of $x: [a, b] \rightarrow \mathbb{R}$ with $V_x[a, b] < \infty$.

Recall: A signed measure μ on $[a, b]$ is the difference $\mu = \mu_+ - \mu_-$ (Hahn-Jordan decomposition) of two mutually singular finite positive measures μ_+ and μ_- . The total variation measure of μ is

$$|\mu| = \mu_+ + \mu_-$$

Hahn-Jordan thm: Given any finite positive measures μ_1 and μ_2 there are μ_+ and μ_- s.t. $\mu = \mu_1 - \mu_2 = \mu_+ - \mu_-$.

Prop. (i) Given a signed measure μ on $[a, b]$, its CDF $x(t) = \mu([a, t])$ is in $BV[a, b]$, càdlàg, and

$$V_x[a, b] = |\mu|([a, b]).$$

(ii) Given $x \in BV[a, b]$ that is right-continuous, there is μ s.t. $x(t) = \mu([a, t])$.

Proof. (i) Let $\mu = \mu_+ - \mu_-$ be a signed measure and x its CDF. Let $a \leq t_0 < t_1 < \dots < t_n \leq b$. Then

$$\sum_{i=1}^n |x(t_i) - x(t_{i-1})| = \sum_{i=1}^n |\mu(t_{i-1}, t_i]| \leq |\mu|(a, b]$$

$$\Rightarrow V_x[a, b] \leq |\mu|(a, b].$$

For the other direction, let $(t_i^{(m)})_{i=0}^{n_m}$ be a sequence of nested partitions with $\max_i |t_i^{(m)} - t_{i-1}^{(m)}| \xrightarrow{m \rightarrow \infty} 0$. It suffices to show that

$$|\mu|(a, b] = \lim_{m \rightarrow \infty} \sum_{i=1}^{n_m} |x(t_i^{(m)}) - x(t_{i-1}^{(m)})|.$$

To see this, define the probability on $(a, b]$ by

$$\nu(dt) = \frac{|\mu|(dt)}{|\mu|(a, b]}.$$

Let $\mathcal{F}_m = \sigma((t_{i-1}^{(m)}, t_i^{(m)}] : 1 \leq i \leq n_m) \subset \mathcal{F}_{m+1}$

$$X = \frac{d\mu}{d|\mu|} = 1_{\text{supp } \mu_+} - 1_{\text{supp } \mu_-}$$

$$X_m = E[X | \mathcal{F}_m].$$

For $s \in (t_{i-1}^{(m)}, t_i^{(m)}]$ then

$$X_m(s) = \frac{\mu((t_{i-1}^{(m)}, t_i^{(m)}])}{|\mu|((t_{i-1}^{(m)}, t_i^{(m)}])} = \frac{x(t_i^{(m)}) - x(t_{i-1}^{(m)})}{|\mu|((t_{i-1}^{(m)}, t_i^{(m)}])}$$

$$\Rightarrow E|X_m| = \frac{1}{\mu([a,b])} \sum_{i=1}^{n_m} |x(t_i^{(m)}) - x(t_{i-1}^{(m)})|$$

and the claim is $E|X_m| \rightarrow 1$. But (X_m) is a bd martingale, so there is Y s.t. $X_m \rightarrow Y$ in L^1

$$\Rightarrow E|X_m| \rightarrow E|Y|.$$

Since $\sigma(\cup \mathcal{F}_n) = \mathcal{B}([a,b])$ in fact $X=Y$ a.s., so

$$E|Y| = E|X| = 1.$$

(ii) Now let $x \in BV[a,b]$. Assume $x(a) = 0$ and set

$$x_{\pm}(t) = \frac{1}{2}(V_x[a,t] \pm x(t)).$$

Claim: x_{\pm} are increasing, i.e. $x_{\pm}(s) \geq x_{\pm}(t)$ if $s > t$.

Let $a \leq t_0 < \dots < t_n \leq t$ be a partition of $[0,t]$,
s.t. $a \leq t_0 < \dots < t_n \leq t \leq s$ is a partition of $[0,s]$.

$$\Rightarrow 2x_{\pm}(s) = V_x[a,s] \pm x(s)$$

$$\geq \underbrace{\sum |x(t_i) - x(t_{i-1})|}_{\geq V_x[a,t] - \varepsilon} + \underbrace{|x(s) - x(t)|}_{\geq \pm x(t)} \pm x(s)$$

$$\Rightarrow x_{\pm}(s) \geq x_{\pm}(t).$$

Claim: If x is right-continuous, so are x_{\pm} .

It suffices to show $v(t) = V_x[0, t]$ is right-continuous.
Let $\beta = v(t+) - v(t)$. Need to show $\beta = 0$. Let $h > 0$ s.t.

$$v(t+h) - v(t) < \beta + \varepsilon$$
$$|x(s) - x(t)| < \varepsilon \quad \text{for } s \in [t, t+h].$$

There is a partition of $[t, t+h]$ s.t.

$$\sum_{i=1}^n |x(t_i) - x(t_{i-1})| \geq \frac{3}{4} V_x[t, t+h] \geq \frac{3}{4} \beta$$
$$\Rightarrow \sum_{i=2}^n |x(t_i) - x(t_{i-1})| \geq \frac{3}{4} \beta - |x(t_1) - x(t_0)| \geq \frac{3}{4} \beta - \varepsilon$$

Since $\beta \leq v(t_1) - v(t_0)$ there is a part of $[t_0, t_1]$ s.t.

$$\sum |x(t'_i) - x(t'_{i-1})| \geq \frac{3}{4} \beta$$

Thus there is a partition of $[t, t+h]$ s.t.

$$V_x[t, t+h] \geq \frac{3}{2} \beta - \varepsilon$$

On the other hand,

$$V_x[t, t+h] \leq \beta + \varepsilon$$

$$\Rightarrow \beta \leq 4\varepsilon \quad \forall \varepsilon > 0 \Rightarrow \beta = 0.$$

Defn. Let $x \in BV[a, b]$ be right-continuous with associated signed measure μ . For $f \in L^1(|\mu|)$ the Lebesgue-Stieltjes integral is defined by

$$\int_s^t f(u) dx(u) = \int_{(s, t]} f(u) \mu(du) \quad (a \leq s < t \leq b)$$

$$\int_s^t f(u) |dx(u)| = \int_{(s, t]} f(u) |\mu|(du)$$

and set $f \circ x(t) = \int_0^t f(u) dx(u)$.

Exercise. Let $x \in BV[a, b]$ be right-continuous. Then

$$\left| \int_a^b f(t) dx(t) \right| \leq \int_a^b |f(t)| |dx(t)|$$

and $f \circ x$ is in BV and right-cont. with signed measure $f(t) dx(t)$ and TV $|f(t)| |dx(t)|$.

Prop. Let $x \in BV$ be right-cont. and f bounded and left-cont. Then for any sequence of partitions $(t_i^{(m)})_{i=1}^{n_m}$ of $[a, b]$ with step size $\rightarrow 0$:

$$\int_a^b f(t) dx(t) = \lim_{m \rightarrow \infty} \sum_{i=1}^{n_m} f(t_{i-1}^{(m)}) (x(t_i) - x(t_{i-1}))$$

$$\int_a^b |f(t)| |dx(t)| = \lim_{m \rightarrow \infty} \sum_{i=1}^{n_m} |f(t_{i-1}^{(m)})| |x(t_i) - x(t_{i-1})|.$$

Proof. Let $f_m(a) = 0$, $f_m(t) = f(t_{i-1}^{(m)})$ if $t \in (t_{i-1}^{(m)}, t_i^{(m)}]$.

$\Rightarrow f(t) = \lim_{m \rightarrow \infty} f_m(t)$ by left-continuity.

$$\begin{aligned} \Rightarrow \sum_{i=1}^{n_m} f(t_{i-1}^{(m)}) (x(t_i^{(m)}) - x(t_{i-1}^{(m)})) &= \int_{(a,b]} f_m(t) dx(t) \\ &\stackrel{\text{DCT}}{\rightarrow} \int f(t) dx(t). \end{aligned}$$

The second claim is similar.

Exercise. Let $x, y \in BV[a, b]$ be right-cont. Then

$$x(b)y(b) - x(a)y(a) = \int_a^b x(t) dy(t) + \int_a^b y(t-) dx(t)$$

1.3. Continuous-time martingales

Defn. Given a filtered probability space $(\Omega, \mathcal{F}, (\mathcal{F}_t), P)$, a martingale is a family of random variables $M_t, t \in \mathbb{T}$ such that

(i) $(t \mapsto M_t) \in D(\mathbb{T}), P$ -a.s. [if \mathbb{T} is continuous]

(ii) $\forall t \geq 0: M_t \in \mathcal{F}_t$ and $E|M_t| < \infty$.

(iii) $\forall 0 \leq s \leq t: E[M_t | \mathcal{F}_s] = M_s, P$ -a.s.

If $(t_k) \subset \mathbb{T}$ then $(M(t_k))_k$ is a discrete-time martingale w.r.t. (\mathcal{F}_{t_k}) . By taking $t_k = 2^{-n}k$ and using right-continuity, most properties extend from discrete to continuous.

Doob's inequality. If (M_t) is a martingale and

$$E_\ell = \left\{ \omega: \sup_{0 \leq s \leq t} |M_s(\omega)| \geq \ell \right\}$$

then

$$P[E_\ell] \leq \frac{1}{\ell} E[1_{E_\ell} |M_t|] \leq \frac{1}{\ell} E[|M_t|]$$

$$P[E_\ell] \leq \frac{1}{\ell^2} E[1_{E_\ell} |M_t|^2] \leq \frac{1}{\ell^2} E[|M_t|^2].$$

Proof. The discrete-time versions of Doob's inequalities imply the statement with E_e replaced by

$$E_e^n = \left\{ \sup_{\substack{0 \leq s \leq t \\ s \in 2^{-n} \mathbb{N}}} |M_s| \geq \ell \right\}.$$

By right-continuity, $P(E_e^n) \uparrow P(E_e)$.

Defn. (X_t) is a submartingale if (i) & (ii) hold,

(iii') $\forall 0 \leq s \leq t: E[X_t | \mathcal{F}_s] \geq X_s$, P -a.s.

and a supermartingale if $E[X_t | \mathcal{F}_s] \leq X_s$.

Exercise. Let (X_t) be a martingale, $\phi: \mathbb{R} \rightarrow \mathbb{R}$ convex, and assume $E[|\phi(X_T)|] < \infty$. Then $\phi(X_t)$ is integrable for any $t \in T$ and a submartingale.

Defn. A stopping time is a random variable τ with values in $[0, \infty]$ s.t.

$$\forall t \geq 0: E_t = \{ \tau \leq t \} \in \mathcal{F}_t.$$

Define $\mathcal{F}_\tau = \{ A \in \mathcal{F} : A \cap E_t \in \mathcal{F}_t \}$.

Exercise. • If τ_1, τ_2 are stopping times, so are $\tau_1 \wedge \tau_2$ and $\tau_1 \vee \tau_2$.

- If $\tau_1 \leq \tau_2$ are stopping times, $\mathcal{F}_{\tau_1} \subseteq \mathcal{F}_{\tau_2}$.
- If f is a measurable function s.t. $f(t) \geq t$ for all t then $f(\tau)$ is a stopping time if τ is one. Thus $\tau_n = 2^{-n} \lceil 2^n \tau \rceil$ is a stopping time s.t. $\tau_n \geq \tau$ and $\tau_n \rightarrow \tau$.

Exercise: Let X be a random variable in L^1 i.e. $E|X| < \infty$. Then the set of random variables $E[X | \mathcal{G}]$ where $\mathcal{G} \subset \mathcal{F}$ ranges over sub- σ -fields is uniformly integrable: $\forall \varepsilon > 0 \exists \lambda > 0$ s.t. $\forall \mathcal{G}$:

$$E[|E[X | \mathcal{G}]| \mathbb{1}_{|E[X | \mathcal{G}]| > \lambda}] \leq \varepsilon.$$

Vitali's convergence theorem: Equivalent:

- $X_j \in L^1$ and $X_j \rightarrow X$ in L^1
- X_j UI and $X_j \rightarrow X$ in probability.

Thm. (OST). Let (M_t) be a martingale and $\tau_1 \leq \tau_2$ be two bounded stopping times. Then $M_{\tau_i} \in L^1$ and

$$E[M_{\tau_2} | \mathcal{F}_{\tau_1}] = M_{\tau_1} \quad \text{a.s.}$$

Proof. Assume $\tau_1, \tau_2 \leq N$. Let $\tau_i^n = 2^{-n} \lceil 2^n \tau_i \rceil \leq N+1$ where N is a deterministic constant. The discrete-time OST implies that

$$M_{\tau_2^n} = E[M_{N+1} | \mathcal{F}_{\tau_2^n}]$$

In particular, $M_{\tau_2^n}$ is UI. Since $M_{\tau_2^n} \rightarrow M_{\tau_2}$ a.s. therefore $M_{\tau_2^n} \rightarrow M_{\tau_2}$ in L^1 and likewise for τ_1 . The discrete OST also implies

$$E[M_{\tau_2^n} | \mathcal{F}_{\tau_1^n}] = M_{\tau_1^n}$$

In particular, for any $A \in \mathcal{F}_{\tau_1} \subseteq \mathcal{F}_{\tau_1^n}$,

$$E[M_{\tau_2^n} 1_A] = E[M_{\tau_1^n} 1_A]$$

Since $M_{\tau_i^n} \rightarrow M_{\tau_i}$ in L^1 the claim follows.

Cor. For any stopping time, $M_{t \wedge \tau}$ is a martingale.
Exercise. Extend to (M_t) a submartingale.

1.4. Semimartingales

Let $(\Omega, \mathcal{F}, (\mathcal{F}_t), P)$ be a filtered probability space.

Example Let (X_n) be an adapted discrete-time process. Define

$$A_n = \sum_{j=1}^n E[X_j - X_{j-1} | \mathcal{F}_{j-1}], \quad A_0 = 0.$$

Then $Y_n = X_n - A_n$ is a martingale: for $m \leq n$,

$$\begin{aligned} E[Y_n | \mathcal{F}_m] &= E[X_n | \mathcal{F}_m] - \sum_{j=1}^n \underbrace{E[E[X_j - X_{j-1} | \mathcal{F}_{j-1}] | \mathcal{F}_m]} \\ &= \begin{cases} E[X_j - X_{j-1} | \mathcal{F}_m] & (j > m) \\ E[X_j - X_{j-1} | \mathcal{F}_{j-1}] & (j \leq m) \end{cases} \\ &= X_m - \sum_{j=1}^m E[X_j - X_{j-1} | \mathcal{F}_{j-1}]. \end{aligned}$$

and $A_n \in \mathcal{F}_{n-1}$. The decomposition $X_n = Y_n + A_n$ with these properties is the Doob or semimartingale decomposition of X_n . The process Y_n is the martingale part of X_n and A_n the compensator.

Exercise: The semimartingale decomposition is unique.

The semimartingale decomposition is more subtle in the continuous case.

Prop. Let (A_t) be a continuous-time martingale. If A is continuous and BV, P-a.s., then A is const.

Proof. Assume $A_0 = 0$ and let A_t^* be the variation of A on $[0, t]$. It suffices to consider $A^*(t) \leq C$. Otherwise, consider $A_{t \wedge \tau}$ where $\tau = \inf\{t: A_t^* \geq C\}$.

Let $t_j = (j/n)T$. Then

$$\begin{aligned} \sum_{j=1}^n |A_{t_j} - A_{t_{j-1}}|^2 &\leq \sup_j |A_{t_j} - A_{t_{j-1}}| \times \sum_j |A_{t_j} - A_{t_{j-1}}| \\ &\leq C \sup_j |A_{t_j} - A_{t_{j-1}}|. \end{aligned}$$

Since $|A_{t_j}| \leq |A_0| + C$ and A is continuous on $[0, T]$, thus uniformly continuous,

$$E\left[\sum_{j=1}^n |A_{t_j} - A_{t_{j-1}}|^2\right] \rightarrow 0.$$

On the other hand, since A is a martingale,

$$E\left[\sum_{j=1}^n |A_{t_j} - A_{t_{j-1}}|^2\right] = E\left[|A_T - A_0|^2\right]$$

Thus $E\left[|A_T - A_0|^2\right] = 0$, i.e. $A_T = A_0$ a.s.

Prop. Let (M_t) be a bounded martingale and (A_t) be an adapted continuous process of finite variation with $A_t^* \leq C_t$ for every t . Then

$$X_t = M_t A_t - M_0 A_0 - \int_0^t M_s dA_s$$

is a martingale.

Rk. Formally, $X_t = \int A_s dM_s$.

Proof. Since A is continuous,

$$\begin{aligned} \int_0^t M_s dA_s &= \lim_{n \rightarrow \infty} \sum_{j=1}^n M_{t_j} (A_{t_j} - A_{t_{j-1}}) \\ &= M_t A_t - M_0 A_0 + \lim_{n \rightarrow \infty} \sum_{j=1}^n A_{t_{j-1}} (M_{t_j} - M_{t_{j-1}}) \\ \Rightarrow X_t &= \lim_{n \rightarrow \infty} \sum_{j=1}^n A_{t_{j-1}} (M_{t_j} - M_{t_{j-1}}) = \lim_{n \rightarrow \infty} X_t^n \end{aligned}$$

Since $A_{t_{j-1}} \in \mathcal{F}_{t_{j-1}}$ and M is a martingale,

$$E[A_{t_{j-1}} (M_{t_j} - M_{t_{j-1}}) | \mathcal{F}_{t_{j-1}}] = 0.$$

and similarly X^n is a martingale.

If M is bd, $X_t^n \rightarrow X_t$ in L^1 . It follows that X is a martingale.

2. Brownian motion

In fact, suffices to have t in dense set.

2.1. Kolmogorov's continuity criterion

Thm. (Kolmogorov's criterion). Assume $(X_t)_{t \in \mathbb{R}^d}$ is a stochastic process s.t. for some $p \geq 1$, $\alpha > 0$,

$$\mathbb{E}[|X_t - X_s|^p] \leq A |t-s|^{d+\alpha}, \quad 0 \leq t, s \leq 1.$$

Then there is a version \tilde{X} of X , i.e. $P(X_t = \tilde{X}_t) = 1 \forall t$, that is continuous, in fact γ -Hölder continuous:

$$P\left[\sup_{t \neq s} \frac{|\tilde{X}_t - \tilde{X}_s|}{|t-s|^\gamma} \geq \lambda\right] \leq C \frac{A}{\lambda^p}$$

provided that $\gamma < \alpha/p$ and with $C = C(\alpha, \gamma, p)$.

The essence is really a real analysis estimate. Formally, the estimate follows taking expectation over the following inequality ("Besov space embedding").

Prop. For $x: \mathbb{R}^d \rightarrow \mathbb{R}^n$ continuous, $p \geq 1$ and $\gamma > 0$,

$$\underbrace{\sup_{\substack{t \neq s \\ t, s \in B}} \frac{|x(t) - x(s)|}{|t-s|^\gamma}}_{[x]_\gamma} \leq C_{\gamma, p} \underbrace{\left(\iint_{B, B} \frac{|x(t) - x(s)|^p}{|t-s|^{\gamma p + 2d}} dt ds \right)^{1/p}}_{[x]_{\gamma + 2d/p, p}} \quad (*)$$

where $B = B_1(0) \subset \mathbb{R}^d$.

Garsia-Rodemich-Rumsey inequality. Let p and ψ be strictly increasing continuous functions on $(0, \infty)$ s.t. $p(0) = 0$ and $\psi(t) \rightarrow \infty$ ($t \rightarrow \infty$). Then if $x : [0, T] \rightarrow \mathbb{R}^n$ is continuous and

$$\int_0^T \int_0^T \psi\left(\frac{|x(t) - x(s)|}{p(t-s)}\right) ds dt \leq B < \infty$$

then

$$|x(t) - x(s)| \leq \int_0^{t-s} \psi^{-1}\left(\frac{\psi B}{u^2}\right) p'(u) du$$

Proof of (f) when $d=1$. Let $\psi(u) = u^p$, $p(u) = u^{\delta+2/p}$,

$$B = B(x) = \int_0^1 \int_0^1 \frac{|x(t) - x(s)|^p}{|t-s|^{\delta p+2}} ds dt.$$

$$\begin{aligned} \Rightarrow |x(t) - x(s)| &\leq C_p \int_0^{t-s} \left(\frac{\psi B}{u^2}\right)^{1/p} u^{\delta+2/p-1} du \\ &= C'_p B^{1/p} \int_0^{t-s} u^{\delta-1} du \leq C_{p,\delta} B^{1/p} |t-s|^\delta \end{aligned}$$

Rk. If x is only in L_{loc} and the RHS is finite, there is \tilde{x} s.t. $\tilde{x} = x$ a.e. s.t. the estimate holds: Set

$$x^\varepsilon(t) = \frac{1}{2\varepsilon} \int_{t-\varepsilon}^{t+\varepsilon} x(t+u) du$$

Then $\{x^\varepsilon\}_\varepsilon$ is bounded. By Arzela-Ascoli there is \tilde{x} s.t. $\|x^\varepsilon - \tilde{x}\|_\infty \rightarrow 0$ along some $\varepsilon \downarrow 0$. But $x^\varepsilon \rightarrow x$ a.e. by Lebesgue diff. thm., so $\tilde{x} = x$ a.e.

Proof of Kolmogorov's criterion. If we already knew that X was continuous, the estimate would follow from the GRR inequality. But we have not even assumed that $t \mapsto X_t$ is Borel measurable!

Let

$$X_t^n = X_{2^{-n}(2^n t)}$$

Then X_t^n is measurable. Set

$$\hat{X}_t(\omega) = \begin{cases} \lim_{n \rightarrow \infty} X_t^n(\omega) & \text{for } (\omega, t) \text{ s.t. the limit exists} \\ 0 & \text{else.} \end{cases}$$

Since $P[|X_t^n - X_t| \geq 2^{-n}] \leq A 2^{+n} (2^{-n})^{1+\beta}$ is summable the limit exist a.s. for every fixed t .

Let

$$Y = \left[\mathbb{I} \hat{X} \mathbb{I} = \left(\int_0^1 \int_0^1 \frac{|\hat{X}_t - \hat{X}_s|^p}{|t-s|^{\alpha p + 2}} ds dt \right)^{1/p} \right]$$

$$\Rightarrow E[Y^p] = A \int_0^1 \int_0^1 |t-s|^{1+\alpha - p\alpha - 2} ds dt < \infty \text{ if } \alpha > \alpha p.$$

Thus, for every ω , \hat{X} coincides for a.e. t with a \tilde{X} s.t.

$$[\hat{X}]_\alpha \leq C Y \quad \text{for all } \omega.$$

It remains to see that $P[X_t = \tilde{X}_t] = 1$ for every t .
 It suffices to show this for $t = 2^{-n}k$. Then

$$P[X_t = \tilde{X}_t] = P[\hat{X}_t = \tilde{X}_t].$$

But $P[|\hat{X}_t - \tilde{X}_t^\varepsilon| > \delta] \lesssim \delta^{-1} \varepsilon^{1+\beta} \rightarrow 0$, so as before
 $\tilde{X}_t^\varepsilon \rightarrow \tilde{X}_t$ a.s. as $\varepsilon \rightarrow 0$, so $\hat{X}_t = \tilde{X}_t$ a.s.

Proof (GRR). Let

$$I(t) = \int_0^t \psi\left(\frac{|x(t) - x(s)|}{p(t-s)}\right) ds$$

so that $\int_0^T I(t) dt = B$. By the mean-value theorem,
 there is $t_0 \in (0, T)$ s.t. $I(t_0) \leq B/T$.

Given t_{n-1} , define d_{n-1} by

$$p(d_{n-1}) = \frac{1}{2} p(t_{n-1}) \quad (\text{so } 0 < d_{n-1} < t_{n-1})$$

and $t_n \in (0, d_{n-1})$ s.t.

$$I(t_n) \leq \frac{2B}{d_{n-1}} \quad \& \quad \psi\left(\frac{|x(t_n) - x(t_{n-1})|}{p(t_n - t_{n-1})}\right) \leq 2 \frac{I(t_{n-1})}{d_{n-1}}.$$

Each can fail on set of measure $< \frac{d_{n-1}}{2}$ of t_n .

Thus both can fail only on a set of meas. $< d_{n-1}$.

Thus

$$t_0 > d_0 > t_1 > d_1 > \dots$$

Then:

$$p(d_{n+1}) = \frac{1}{2} p(t_{n+1}) \leq \frac{1}{2} p(d_n) \Rightarrow d_n, t_n \downarrow 0.$$

Also:

$$\begin{aligned} p(t_n - t_{n+1}) &\leq p(t_n) = 2p(d_n) = 4(p(d_n) - \frac{1}{2}p(d_n)) \\ &\leq 4(p(d_n) - p(d_{n+1})) \end{aligned}$$

$$\begin{aligned} \Rightarrow |x(t_{n+1}) - x(t_n)| &\leq p(t_n - t_{n+1}) \Psi^{-1}\left(\frac{2I(t_n)}{d_n}\right) \\ &\leq 4(p(d_n) - p(d_{n+1})) \Psi^{-1}\left(\frac{4B}{d_n d_{n-1}}\right) \\ &\leq 4 \int_{d_{n+1}}^{d_n} \Psi^{-1}\left(\frac{4B}{u^2}\right) p(du) \leq \frac{4B}{d_n^2} \end{aligned}$$

$$\Rightarrow |x(t_0) - x(0)| \leq 4 \int_0^T \Psi^{-1}\left(\frac{4B}{u^2}\right) p(du)$$

By an analogous argument,

$$|x(T) - x(t_0)| \leq 4 \int_0^T \Psi^{-1}\left(\frac{4B}{u^2}\right) p(du)$$

$$\Rightarrow |x(T) - x(0)| \leq 8 \int_0^T \Psi^{-1}\left(\frac{4B}{u^2}\right) p(du).$$

Given $0 \leq s < t \leq T$ set

$$\bar{x}(u) = x\left(s + \frac{(t-s)}{T} u\right)$$

$$\bar{p}(u) = p\left(\frac{(t-s)}{T} u\right)$$

The argument above with $B \rightsquigarrow \left(\frac{T}{t-s}\right)^2 B = \bar{B}$ gives

$$|x(t) - x(s)| \leq 8 \int_0^T \psi^{-1}\left(\frac{4B}{u^2}\right) \bar{p}(du)$$

$$= 8 \int_0^{t-s} \psi^{-1}\left(\frac{4\bar{B}}{u^2}\right) p(du).$$

2.2. Definition of Brownian motion

Defn. A stochastic process $(B_t)_{t \geq 0}$ is called a Brownian motion or Wiener process if it is continuous and:

(i) $\forall t < s: B_t - B_s \sim N(0, t-s)$.

(ii) $\forall t < s: B_t - B_s$ is independent of $\sigma(B_u: u \leq s)$.

Thm. There exists a unique probability measure on $C[0, \infty)$ such that the canonical process is a Wiener process. This is the Wiener measure.

Lemma. If (i) holds for a stochastic process on some probability space then there is a version that is in fact Hölder continuous for any Hölder exponent strictly less than $1/2$.

Proof. Since $W_t - W_s \sim N(0, t-s)$,

$$E[|W_t - W_s|^p] \leq C_p |t-s|^{p/2} \text{ for any } p \geq 1.$$

Thus Kolmogorov applies $\gamma = \frac{1}{2} - \varepsilon$ when $p \geq \frac{1}{\varepsilon}$:

$$E[|W_t - W_s|^p] \leq C_p |t-s|^{p/2} \leq C_p |t-s|^{\gamma p + 1}$$

Proof (construction of Wiener measure). Given any probability space (Ω, \mathcal{F}, P) on which a Brownian motion (B_t) is defined, consider the map

$$I: \Omega \rightarrow C[0, \infty), \quad \omega \mapsto (t \mapsto B_t(\omega)).$$

This map is measurable. Thus $P \circ I^{-1}$ defines a Borel measure on $C[0, \infty)$.

Using a monotone class argument one checks it is uniquely defined by its fin-dim. distributions.

Prop. Brownian motion is not in $C^{1/2}$:

$$P \left[\sup_{\substack{t \neq s \\ 0 < t, s < 1}} \frac{|X_t - X_s|}{|t - s|^{1/2}} = \infty \right] = 1.$$

Proof. For each $t > s$, $(B_t - B_s)/|t - s|^{1/2}$ is $\sim N(0, 1)$.

$$\sup_{t \neq s} \frac{|X_t - X_s|}{|t - s|^{1/2}} \geq \sup_{0 \leq k < 2^n} \frac{|X_{2^{-n}(k+1)} - X_{2^{-n}k}|}{2^{-n/2}}$$

$$\Rightarrow P \left[\sup \frac{|X_t - X_s|}{|t - s|^{1/2}} > \lambda \right] \leq 1 - P[Z < \lambda]^{2^n} \rightarrow 0.$$

where $Z \sim N(0, 1)$.

There are many ways of constructing (B_t) .

Defn. Let (Ω, \mathcal{F}, P) be a probability space. Then $S \subset L^2(\Omega, \mathcal{F}, P)$ is a Gaussian space if S is a closed linear subspace and any X is a Gaussian random variable.

Example. Let (X_i) be i.i.d $\mathcal{N}(0, 1)$ on some probability space. Then $\text{span}\{X_i\}$ is a Gaussian space. The X_i are an orthonormal system:

$$E[X_i X_j] = \delta_{ij}$$

Note that limits of Gaussian random variables are Gaussian.

Prop. Let H be a separable Hilbert space and S as in the example. Then there is an isometry $I: H \rightarrow S$. Thus:

- $\forall f \in H: I(f) \sim \mathcal{N}(0, \|f\|_H^2)$
- $\forall f, g \in H: E I(f) I(g) = (f, g)_H$

In fact, $I(\alpha f + \beta g) = \alpha I(f) + \beta I(g)$ a.s.

Proof. Let (e_i) be an orthonormal basis for H . Set

$$I(f) = \sum_{i=1}^{\infty} (f, e_i) X_i.$$

The limit exists in L^2 and a.s. since

$$k \mapsto \sum_{i=1}^k (f, e_i) X_i$$

is a martingale bounded in L^2 .

Defn. Let \dot{W} be an isometry from $L^2(\mathbb{R}_+)$ into some Gaussian space. Then \dot{W} is called white noise.

For $t \geq 0$, define

$$W_t = \dot{W}(1_{[0,t]})$$

Rk. Would like to think of $\dot{W}(f)$ as $\int f dW$ but W has infinite variation almost surely.

Exercise. (i) For $A \subset \mathbb{R}_+$ Borel, $|A| < \infty$, $\dot{W}(A) \sim N(0, |A|)$

(ii) For $A, B \subset \mathbb{R}_+$ with $A \cap B = \emptyset$, $\dot{W}(A)$ and $\dot{W}(B)$ are independent.

(iii) For $A = \bigcup_{i=1}^{\infty} A_i$ with A_i disjoint,
 $\dot{W}(A) = \sum_{i=1}^{\infty} \dot{W}(A_i)$ in L^2 and a.s.

W looks like a random measure but it is not.

Prop. For any t_1, \dots, t_n the vector $(W_{t_i})_{i=1}^n$ is jointly Gaussian with covariance

$$E[W_t W_s] = s \wedge t \quad \text{for } s, t \geq 0.$$

Moreover, $W_0 = 0$ a.s. and

$W_t - W_s$ is independent of $\sigma(W_u : u < s)$

$$W_t - W_s \sim N(0, t-s).$$

Exercise: Let B be a Brownian motion. Then

- $-B$ is a Brownian motion (symmetry)
- $\frac{1}{\lambda} B_{\lambda^2 t}$ is a Brownian motion (scale invariance)
- $B_{t+s} - B_s$ is a Brownian motion that is indep. of $(B_u : u \leq s)$ (Markov prop.).

Exercise. Understand how Brownian motion is the Gaussian measure with covariance $(-\Delta)^{-1}$ on $[0, \infty)$ with 0-boundary condition at 0.

Defn. A stochastic process taking values in \mathbb{R}^d is a d -dim. Brownian motion if each component is a Brownian motion.

2.3. Heat equation and Feynman-Kac formula

Let $p_t(x,y) = \frac{1}{\sqrt{2\pi t}} e^{-\frac{|x-y|^2}{2t}}$ be the heat kernel on \mathbb{R} .

Defn. The semigroup of Brownian motion is

$$T_t f(x) = E_x[f(B_t)]$$

where E_x is the expectation of a BM starting at $x \in \mathbb{R}$.

Fact. Since $B_t - x \sim N(0,t)$,

$$\begin{aligned} T_t f(x) &= \int p_t(x,y) f(y) dy = \int f(x+y) p_t(0,y) dy \\ &= \int f(x+y\sqrt{t}) p_t(0,y) dy. \end{aligned}$$

Defn. The infinitesimal generator is

$$Lf(x) = \lim_{t \downarrow 0} \frac{T_t f(x) - f(x)}{t}$$

for f such that the limit exists.

Fact. For $f \in C^2$ the limit exists and

$$Lf = \frac{1}{2} f''.$$

Proof. If $f \in C^3$,

$$\lim_{t \rightarrow 0} \int \frac{f(x+y\sqrt{t}) - f(x)}{t} p_1(0,y) dy$$

$$= \lim_{t \rightarrow 0} \int \frac{f'(x)y\sqrt{t} + \frac{1}{2}f''(x)y^2t + t^{3/2}e_t(y)}{t} p_1(0,y) dy$$

with $|e_t(y)| \leq C|y|^2$. Thus

$$Lf(x) = \frac{1}{2}f''(x).$$

If $f \in C^2$, one still has $t^{3/2}|e_t(y)| = o(t)$.

Thus $u(t,x) = T_t f(x)$ satisfies the heat equation

$$\frac{\partial u}{\partial t} = \frac{1}{2} \frac{\partial^2 u}{\partial x^2}, \quad u(t,x) \rightarrow f(x) \quad (t \rightarrow 0).$$

Exercise. Let $X_t = x + \sigma B_t + mt$ be Brownian motion with variance σ^2 and drift m . Show

$$Lf(x) = \frac{\sigma^2}{2} f''(x) + m f'(x).$$

Fact. Given $(\Omega, \mathcal{F}, (\mathcal{F}_t), P)$, a continuous adapted process (B_t) is a Brownian motion iff

$$E[f(B_t) | \mathcal{F}_s] = T_{t-s} f(B_s) \quad \forall f \text{ continuous.}$$

Feynman-Kac formula. Let V be bounded and continuous. Then

$$u(t, x) = E_x \left[\exp\left(\int_0^t V(B_s) ds\right) f(X_t) \right]$$

satisfies

$$\frac{\partial u}{\partial t} = \frac{1}{2} \frac{\partial^2 u}{\partial x^2} + Vu, \quad u(0, x) = f(x).$$

Proof. Expand exponential:

$$\begin{aligned} u(t, x) &= \sum_{n=0}^{\infty} \frac{1}{n!} E_x \left[\int_0^t \dots \int_0^t V(B_{t_1}) \dots V(B_{t_n}) f(X_t) dt_1 \dots dt_n \right] \\ &= \sum_{n=0}^{\infty} \tilde{E}_x \left[\int_{0 < t_1 < \dots < t_n < t} V(B_{t_1}) \dots V(B_{t_n}) f(X_t) dt_1 \dots dt_n \right] \\ &= \sum_{n=0}^{\infty} E_x \left[\int_{0 < t_n < \dots < t_1 < t} V(B_{t_1}) \dots V(B_{t_n}) f(X_t) dt_1 \dots dt_n \right] \\ &= \sum_{n=0}^{\infty} \int_{0 < t_n < \dots < t_1 < t} T_{t-t_1} V T_{t_1-t_2} \dots V T_{t_{n-1}-t_n} V T_{t_n} f dt_1 \dots dt_n \end{aligned}$$

Differentiate term by term.

Example: Arcsine law.

$$\text{Let } \bar{z}_t = \frac{1}{t} \int_0^t 1_{B_s > 0} ds.$$

$$\Rightarrow E[\bar{z}_t] = \frac{1}{2}.$$

What is the distribution of \bar{z}_t ?

$$\bar{z}_t = \int_0^1 1_{B_{ts} > 0} ds \stackrel{(d)}{=} \int_0^1 1_{\sqrt{t}B_s > 0} ds = \int_0^1 1_{B_s > 0} ds$$

$$\Rightarrow \bar{z}_t \sim \bar{z}_1.$$

Prop. $P[\bar{z}_1 \leq x] = \int_0^x \frac{1}{\pi \sqrt{y(1-y)}} dy = \frac{2}{\pi} \arcsin(\sqrt{x}).$

Proof. Let $V(x) = 1_{x > 0}$. This V is not cont. and a more careful argument would involve a limiting argument to apply Feynman-Kac. Anyway,

$$u(t, x) = E_x \left[\exp\left(-\sigma \int_0^t V(B_s) ds\right) \right]$$

satisfies

$$\frac{\partial u}{\partial t} = \frac{1}{2} \frac{\partial^2 u}{\partial x^2} - \sigma V u, \quad u(0, x) = 1.$$

Let $g(\lambda, x) = \int_0^\infty e^{-\lambda t} u(t, x) dt$. Then

$$\lambda g + \sigma V g - \frac{1}{2} \frac{\partial^2 g}{\partial x^2} = 0.$$

Thus

$$(\lambda + \sigma)g - \frac{1}{2}g'' = 1 \quad (x > 0)$$

$$\lambda g - \frac{1}{2}g'' = 0 \quad (x < 0)$$

and g should be bounded at $\pm\infty$ and continuously differentiable at 0. The solution is

$$g(\lambda, x) = \frac{1}{\sqrt{\lambda}} \frac{1}{\sqrt{\lambda + \sigma}}.$$

Thus

$$\begin{aligned} & E\left[\int_0^{\infty} \exp(-\lambda t - \sigma \int_0^t V(B_s) ds) dt\right] \\ &= E\left[\int_0^{\infty} \exp(-\lambda t - \sigma t \xi_t) dt\right] \\ &= E\left[\int_0^{\infty} \exp(-\lambda t - \sigma t \xi_1) dt\right] \\ &= E\left[\frac{1}{\lambda + \sigma \xi_1}\right] = \frac{1}{\sqrt{\lambda}} \frac{1}{\sqrt{\lambda + \sigma}}. \end{aligned}$$

Thus $E\left[\frac{1}{\lambda + \sigma \xi_1}\right] = \frac{1}{\sqrt{\lambda + \sigma}}$. Differentiating:

$$E[\xi_1^n] = \frac{1}{\Gamma} \int_0^1 x^{n-\frac{1}{2}} (1-x)^{-\frac{1}{2}} dx.$$

Since the support is bd. moments characterize the distribution.

2.4. Strong Markov and reflection property

Markov property. Let $t > s$. Then

$$\Leftrightarrow \forall A: P[B_t \in A | \mathcal{F}_s] = P_{t-s}(B_s, A) = \frac{1}{\sqrt{2\pi(t-s)}} \int_A e^{-\frac{(y-B_s)^2}{2(t-s)}} dy.$$
$$\Leftrightarrow \forall f \text{ bd: } P[f(B_t) | \mathcal{F}_s] = P_{t-s} f(B_s) = \int P_{t-s}(B_s, y) f(y) dy.$$

Proof. Immediate from independent increments.

Strong Markov property. Let τ be a stopping time. Then $B_{\tau+t} - B_\tau$ is again a BM independent of \mathcal{F}_τ conditional on $\tau < \infty$.

Proof. Let $\tau_n = 2^{-n} \lceil 2^n \tau \rceil$ and $t_k = 2^{-n} k$.

Given $A \in \mathcal{F}_{\tau+t}$ set $A_k = A \cap \{\tau_n = t_k\} \in \mathcal{F}_{t_k+t}$. Then

$$\begin{aligned} E[f(B_{\tau_n+t}) 1_{A_k}] &= E[f(B_{t_k+t}) 1_{A_k}] \\ &= E[T_{t-s} f(B_{t_k+s}) 1_{A_k}] \\ &= E[T_{t-s} f(B_{\tau_n+s}) 1_{A_k}] \end{aligned}$$

$$\stackrel{\text{M}}{\Rightarrow} E[f(B_{\tau_n+t}) 1_{A \cap \{\tau < \infty\}}] = E[T_{t-s} f(B_{\tau_n+s}) 1_{A \cap \{\tau < \infty\}}]$$

For f bounded, can take $n \rightarrow \infty$:

$$E[f(B_{\tau+t}) 1_{A \cap \{\tau < \infty\}}] = E[T_{t-s} f(B_{\tau+s}) 1_{A \cap \{\tau < \infty\}}]$$

$$\Rightarrow E[f(B_{\tau+t}) | \mathcal{F}_{\tau+s}] = E[T_{t+s} f(B_{\tau+s})] \text{ a.s. on } \tau < \infty$$

Rk. Let $\theta_s : C[0, \infty) \rightarrow C[0, \infty)$, $\theta_s \omega(t) = \omega(t+s)$.

Then for fin. stopping times

$$P_x[\theta_\tau \omega \in A | \mathcal{F}_\tau] = P_{x(\tau)}[A], \text{ } P_x\text{-a.s.}$$

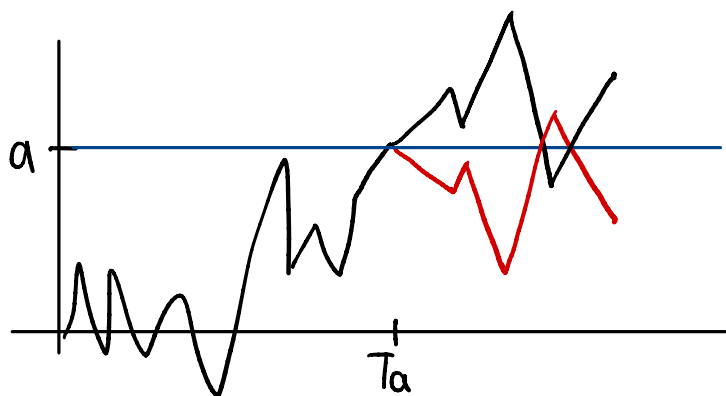
For a stopping time that is not a.s. finite,

$$P_x[\theta_\tau \omega \in A | \mathcal{F}_\tau] = P_{x(\tau)}[A], \text{ } P_x\text{-a.s. on } \tau < \infty.$$

Reflection principle. Let $a > 0$, $T_a = \inf\{t \geq 0 : B_t = a\}$, and

$$X_t = \begin{cases} B_t & (t \leq T_a) \\ 2a - B_t & (t \geq T_a) \end{cases}$$

Then (X_t) is also a Brownian motion.



Proof. On the event $T_a < \infty$ set

$$\tilde{B}_t = B_{T_a+t} - B_{T_a}.$$

By the strong Markov property, \tilde{B} is a BM cond. on $T_a < \infty$ indep. of \mathcal{F}_{T_a} . Same for $-\tilde{B}$.

$$\Rightarrow B_t = B_{T_a \wedge t} + \tilde{B}_{(t-T_a)^+} \mathbb{1}_{T_a < \infty}$$

$$X_t = B_{T_a \wedge t} - \tilde{B}_{(t-T_a)^+} \mathbb{1}_{T_a < \infty}.$$

Thus B and X have the same distribution.

Cor. For $a \geq 0$ and $b \leq a$,

$$P\left[\sup_{0 \leq s \leq t} B_s \geq a, B_t \leq b\right] = P[B_t \geq 2a - b]$$

In particular, $\sup_{0 \leq s \leq t} B_s \sim |B_t|$ and

$$P\left[\sup_{0 \leq s \leq t} |B_s| \geq a\right] \leq 4 P[B_t \geq a]$$

Proof. For the first claim notice:

$$\text{LHS} = P[T_a \leq t, B_t \leq b]$$

$$\text{RHS} = P[2a - B_t \leq b]$$

$$= P[T_a \leq t, X_t \leq b] + \underbrace{P[T_a > t, B_t \geq 2a - b]}_0$$

For the second one:

$$P\left[\sup B_s \geq a\right] = \underbrace{P\left[\sup B_s \geq a, B_t \geq a\right]}_{P[B_t \geq a] \text{ trivially}}$$

$$+ \underbrace{P\left[\sup B_s \geq a, B_t \leq a\right]}_{P[B_t \geq a] \text{ by reflection}}$$

$$= 2 P[B_t \geq a]$$

$$= P[|B_t| \geq a]$$